

Lecture 23: Few Applications (BLR Test, LHL)

BLR Linearity Testing I

- We shall consider the following definition of linear functions

Definition (Linear Function)

Let $f: \{0, 1\}^n \rightarrow \{+1, -1\}$ be a boolean function. If $f(x) \cdot f(y) = f(x + y)$, for all $x, y \in \{0, 1\}^n$, then the function f is a linear function.

- Note that χ_S is a linear function for all $S \in \{0, 1\}^n$. In fact, $\{\chi_0, \dots, \chi_{N-1}\}$ is the set of all linear functions.
- Suppose a function f is provided to us as an oracle. We are interested in testing whether it is close-to some linear function. That is, does there exist S such that f and χ_S agree on a large number of inputs, i.e., $f(x) = \chi_S(x)$ for a large fraction of $x \in \{0, 1\}^n$

BLR Linearity Testing II

- Blum–Luby–Rubinfeld provided an algorithm to correctly test this property using only two queries to the f -oracle. This algorithm is known as the BLR linearity testing algorithm
- Here is the pseudo-code of the algorithm

BLR ^{f} :

- Pick random $x, y \in \{0, 1\}^n$ and query f to obtain $u = f(x)$, $v = f(y)$, and $w = f(x + y)$
 - Output true if $u \cdot v == w$
- So, the algorithm is simple. Let us analyze the performance of this algorithm
 - We want to claim that “if the algorithm returns true with high probability then the function f agrees with some χ_S with high probability”

- We make the following claim

Lemma

The probability that our algorithm outputs true is

$$\frac{1 + \sum_{S \in \{0,1\}^n} \hat{f}(S)^3}{2}$$

Proof Outline.

- Note that the algorithm says true when $f(x) \cdot f(y) = f(x+y)$. That is, $f(x) \cdot f(y) \cdot f(x+y) = 1$, because the range of f is $\{+1, -1\}$.
- And, similarly, our algorithm says false when $f(x) \cdot f(y) \cdot f(x+y) = -1$.

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- Therefore, we can conclude that

$$\frac{1}{N^2} \sum_{x,y \in \{0,1\}^n} f(x)f(y)f(x+y) = p - (1-p),$$

where p is the probability that our algorithm says true

- So, to prove the lemma, it suffices to prove that

$$\frac{1}{N^2} \sum_{x,y \in \{0,1\}^n} f(x)f(y)f(x+y) = \sum_{S \in \{0,1\}^n} \hat{f}(S)^3$$

- Let us prove this

$$\begin{aligned}\frac{1}{N^2} \sum_{x,y \in \{0,1\}^n} f(x)f(y)f(x+y) &= \frac{1}{N} \sum_{z \in \{0,1\}^n} \left(\frac{1}{N} \sum_{x \in \{0,1\}^n} f(x)f(z-x) \right) f(z) \\ &= \frac{1}{N} \sum_{z \in \{0,1\}^n} (f * f)(z) \cdot f(z) \\ &= \langle f * f, f \rangle \\ &= \sum_{S \in \{0,1\}^n} \widehat{(f * f)}(S) \cdot \widehat{f}(S) \\ &= \sum_{S \in \{0,1\}^n} \widehat{f}(S)^3\end{aligned}$$

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- Okay, back to our main proof now. Suppose p is the probability that our algorithm outputs true. If $p \geq 1 - \varepsilon$, then, from the lemma above, we have

$$\sum_{S \in \{0,1\}^n} \widehat{f}(S)^3 \geq 1 - 2\varepsilon$$

- Note that Parseval's identity on f implies that

$$\sum_{S \in \{0,1\}^n} \widehat{f}(S)^2 = \langle f, f \rangle = 1,$$

because the range of f is $\{+1, -1\}$

- So, we are given two guarantees

$$\sum_{S \in \{0,1\}^n} \widehat{f}(S)^2 = 1$$

$$\sum_{S \in \{0,1\}^n} \widehat{f}(S)^3 \geq 1 - 2\epsilon$$

We need to prove that $\max_{S \in \{0,1\}^n} \widehat{f}(S)$ is close to 1

- We prove the following result

Lemma

If $\sum_{S \in \{0,1\}^n} \widehat{f}(S)^2 = 1$ and $\sum_{S \in \{0,1\}^n} \widehat{f}(S)^3 \geq 1 - 2\epsilon$ then we have $\max_{S \in \{0,1\}^n} \widehat{f}(S) \geq 1 - 2\epsilon$.

Proof Outline.

$$\begin{aligned}\max_{S \in \{0,1\}^n} \widehat{f}(S) &= \left(\max_{S \in \{0,1\}^n} \widehat{f}(S) \right) \left(\sum_{S \in \{0,1\}^n} \widehat{f}(S)^2 \right) \\ &\geq \sum_{S \in \{0,1\}^n} \widehat{f}(S)^3 \\ &\geq 1 - 2\epsilon\end{aligned}$$

- So, let us recall what we have proven. If the algorithm outputs true with probability $\geq (1 - \epsilon)$, then there exists S such that $\widehat{f}(S) \geq 1 - 2\epsilon$.
- Recall that if q is the probability that f and χ_S agree then we have $\langle f, \chi_S \rangle = q - (1 - q)$. So, $q \geq 1 - \epsilon$, because $\langle f, \chi_S \rangle = \widehat{f}(S)$.

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- Thus, we conclude that if the algorithm outputs true with probability $p \geq (1 - \epsilon)$ then f agrees with some χ_S with probability $q \geq p \geq (1 - \epsilon)$.

Left-over Hash Lemma I

- We need to introduce the definition of a family of universal hash functions.

Definition (Universal Hash Function Family)

Let $\mathcal{H} = \{h_1, \dots, h_\alpha\}$ be a set of $\{0, 1\}^n \rightarrow \{0, 1\}^m$ functions such that for any distinct $x, x' \in \{0, 1\}^n$ we have

$$\mathbb{P} \left[h(x) = h(x') : h \stackrel{s}{\leftarrow} \mathcal{H} \right] \leq \frac{1}{2^m}$$

- Recall that \mathbb{X} has min-entropy k if $\mathbb{P}[\mathbb{X} = x] \leq 2^{-k}$ for any x in the sample space

Left-over Hash Lemma II

- Left-over Hash Lemma (LHL) states the following. The statistical distance between the distributions $(\mathbb{H}(\mathbb{X}), \mathbb{H})$ and (\mathbb{U}, \mathbb{H}) is small, where \mathbb{U} is a uniform distribution over $\{0, 1\}^m$ and \mathbb{H} is the uniform distribution over \mathcal{H} . Formally, it states the following

Lemma (LHL)

Let \mathbb{H} be a uniform distribution over \mathcal{H} , a universal hash function family $\{0, 1\}^n \rightarrow \{0, 1\}^m$, and \mathbb{X} is a random variable over $\{0, 1\}^n$. Then, the following holds

$$\text{SD}((\mathbb{H}(\mathbb{X}), \mathbb{H}), (\mathbb{U}, \mathbb{H})) \leq \frac{1}{2} \sqrt{\frac{2^m}{2^{\text{H}_\infty(\mathbb{X})}}}$$

Proof Outline.

Left-over Hash Lemma III

- We begin with some simplification

$$\begin{aligned} \text{SD}((\mathbb{H}(\mathbb{X}), \mathbb{H}), (\mathbb{U}, \mathbb{H})) &= \mathbb{E} \left[\text{SD}(h(\mathbb{X}), \mathbb{U}) : h \sim \mathbb{H} \right] \\ &\leq \mathbb{E} \left[\frac{M}{2} \sqrt{\sum_{S \in \{0,1\}^m} \widehat{h(\mathbb{X})}(S)^2 - \widehat{h(\mathbb{X})}(0)^2} : h \sim \mathbb{H} \right] \\ &\leq \frac{M}{2} \sqrt{\mathbb{E} \left[\sum_{S \in \{0,1\}^m} \widehat{h(\mathbb{X})}(S)^2 - \frac{1}{M^2} : h \sim \mathbb{H} \right]}, && \text{Jensen} \\ &\leq \frac{M}{2} \sqrt{\mathbb{E} \left[\sum_{S \in \{0,1\}^m} \widehat{h(\mathbb{X})}(S)^2 : h \sim \mathbb{H} \right] - \frac{1}{M^2}} \\ &= \frac{M}{2} \sqrt{\mathbb{E} \left[\langle h(\mathbb{X}), h(\mathbb{X}) \rangle : h \sim \mathbb{H} \right] - \frac{1}{M^2}}, && \text{Parseval} \\ &= \frac{1}{2} \sqrt{M \mathbb{E} [\text{col}(h(\mathbb{X})) : h \sim \mathbb{H}] - 1} \end{aligned}$$

Left-over Hash Lemma IV

- So, we need to estimate $\mathbb{E} [\mathbf{1}_{\{h(x)=h(x')\}} : x \sim \mathbb{X}, x' \sim \mathbb{X}, h \sim \mathbb{H}]$. Note that this is equivalent to the probability that we sample two independent samples $x \sim \mathbb{X}$ and $x' \sim \mathbb{X}$ and it turns out that $h(x) = h(x')$, for $h \sim \mathbb{H}$. That is, the following expression

$$\mathbb{E} [\mathbf{1}_{\{h(x)=h(x')\}} : x \sim \mathbb{X}, x' \sim \mathbb{X}, h \sim \mathbb{H}]$$

- Note that if $x = x'$, then we shall definitely have $h(x) = h(x')$ irrespective of the value of h .
- If $x \neq x'$ then the probability that $h(x) = h(x')$ is $\leq \frac{1}{M}$, for a random $h \sim \mathbb{H}$
- To use these two observations, we proceed formally as follows. We write

$$\mathbf{1}_{\{h(x)=h(x')\}} = \mathbf{1}_{\{x=x'\}} + \mathbf{1}_{\{(x \neq x') \wedge (h(x)=h(x'))\}}$$

So, we have

$$\mathbb{E} [\mathbf{1}_{\{h(x)=h(x')\}}] = \mathbb{E} [\mathbf{1}_{\{x=x'\}}] + \mathbb{E} [\mathbf{1}_{\{(x \neq x') \wedge (h(x)=h(x'))\}}]$$

Left-over Hash Lemma V

- Let p be the collision probability of the random variable \mathbb{X} . We know that $p \leq \frac{1}{K}$, where k is the min-entropy of \mathbb{X} . So, we have $\mathbb{E} \left[\mathbf{1}_{\{x=x'\}} \right] = p$.
- And, by universal hash function family guarantee of \mathcal{H} , we have

$$\mathbb{E} \left[\mathbf{1}_{\{(x \neq x') \wedge (h(x)=h(x'))\}} \right] \leq (1 - p) \frac{1}{M}$$

- So, we have

$$\mathbb{E} [\text{col}(h(\mathbb{X})) : h \sim \mathbb{H}] \leq p + \frac{(1 - p)}{M} < \frac{1}{K} + \frac{1}{M}$$

- Now, going back to our original inequality

$$\begin{aligned} \text{SD} ((\mathbb{H}(\mathbb{X}), \mathbb{H}), (\mathbb{U}, \mathbb{H})) &\leq \frac{1}{2} \sqrt{M \mathbb{E} [\text{col}(h(\mathbb{X})) : h \sim \mathbb{H}] - 1} \\ &< \frac{1}{2} \sqrt{\frac{M}{K}} \end{aligned}$$